Crystallization Learning with the Delaunay Triangulation

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Abstract

Based on the Delaunay triangulation, we propose the crystallization learning to estimate the conditional expectation function in the framework of nonparametric regression. By conducting the crystallization search for the Delaunay simplices closest to the target point in a hierarchical way, the crystallization learning estimates the conditional expectation of the response by fitting a linear model to the data points of the constructed Delaunay simplices. The local feature of crystallization circumvents the computational difficulty of the Delaunay triangulation of the entire feature space, which instead focuses only on the neighbor of the target point and thus greatly expedites the estimation for high-dimensional cases. Because the volumes of Delaunay simplices are adaptive to the density of data points, our method computes neighbor data points uniformly in all directions and thus is more robust to the local geometric structure of the data than existing nonparametric regression methods. We prove the asymptotic properties of the crystallization learning and conduct numerical experiments on both synthetic and real data to demonstrate the advantages of our method in estimation of the conditional expectation function and prediction of the response.

1. Introduction

Consider a regression model,

\[ y_i = E(Y|x_i) + \epsilon_i, \quad i = 1, \ldots, n, \]

where \( x_i \) is a \( d \)-dimensional feature point in the Euclidean space \( \mathbb{R}^d \) (\( n > d \)), \( y_i \) is the observed response and \( \epsilon_1, \ldots, \epsilon_n \in \mathbb{R} \) are independent and identically distributed (i.i.d.) random errors with \( E(\epsilon_i) = 0 \) and \( E(\epsilon_i^2) < \infty \). Nonparametric regression is a collection of methods for estimating the conditional expectation function \( E(Y|z) (z \in \mathbb{R}^d) \) without rigid assumptions on its shape. Recent decades have witnessed extensive research in the field of nonparametric regression, including nearest-neighbor regression (Nadaraya, 1964; Watson, 1964; Cover & Hart, 1967; Benedetti, 1977; Stone, 1977; Altman, 1992), kernel regression (Priestley & Chao, 1972; Hardle & Gasser, 1984; Hein, 2009) and local linear regression (Cleveland, 1979; Cleveland & Devlin, 1988; Fan & Gijbels, 2018). Although the consistency of these methods has been shown under mild conditions, their finite sample performances are sensitive to the local geometric structure of observed feature points. As these methods only consider the distances from the target point \( z \) to observed feature points \( x_1, \ldots, x_n \) in computing the neighbor data points or assigning weights, it is likely that the directions from \( z \) to its neighbors are not uniformly distributed, especially when \( z \) is close to the boundary of the convex hull of feature points or jump points of the feature data density. As a result, the (weighted) mean of neighbor data points may be far from the target point \( z \), leading to large bias in estimating the conditional expectation \( E(Y|z) \).

By incorporating the Delaunay triangulation (Delaunay, 1934) into the framework of nonparametric regression, we propose the crystallization learning which mimics the crystallization process in thermodynamics and circumvents the curse-of-dimensionality issue in the Delaunay triangulation. Based on the DELAUNAYSPARSE algorithm (Chang et al., 2020) which locally constructs the Delaunay simplex \( S(z) \) containing \( z \), we develop the crystallization search for the Delaunay simplices closest to \( S(z) \) and estimate \( E(Y|z) \) by fitting a local linear model to the data points of the obtained Delaunay simplices. Via experiments on synthetic and real data, our method is shown to outperform the existing ones in estimating the conditional expectation function and predicting the response.

2. Methodology

2.1. Delaunay Interpolation

Let \( X \) be a set of \( n \) feature points \( x_1, \ldots, x_n \) in the Euclidean space \( \mathbb{R}^d \) (\( n > d \)). A \( d \)-dimensional triangulation of \( X \), \( \mathcal{T}(X) \), is a mesh of \( d \)-simplices \( \{S_1, \ldots, S_m\} \) satisfying:

1. For \( j = 1, \ldots, m \), the set of \( d + 1 \) vertices of simplex \( S_j \), denoted as \( \mathbb{V}(S_j) \), is a subset of \( X \) and does not lie

in any affine hyperplane of \( \mathbb{R}^d \).

2. For any \( j \neq k \), simplices \( S_j \) and \( S_k \) are disjoint except on their shared boundaries \( S_j \cap S_k \).

3. The union \( S_1 \cup \cdots \cup S_m \) is the convex hull of \( X \), denoted as \( \mathcal{H}(X) \).

Since the \( d \)-simplices \( S_1, \ldots, S_m \) of the triangulation \( T(X) \) fully cover the convex hull \( \mathcal{H}(X) \), for each internal point \( z \in \mathcal{H}(X) \), there exists a simplex \( S(z) \in T(X) \) such that \( z \in S(z) \). Let \( i_1(z), \ldots, i_{d+1}(z) \) denote the indices corresponding to the data points of \( S(z) \), and then there exist \( d+1 \) values \( \gamma_1, \ldots, \gamma_{d+1} \in [0,1] \) such that \( \sum_{k=1}^{d+1} \gamma_k x_{i_k(z)} = z \) and \( \sum_{k=1}^{d+1} \gamma_k = 1 \). Among all triangulations, the Delaunay triangulation is widely used for multivariate interpolation (de Berg et al., 2008) due to its smoothness property. Let \( B_j \) be the open ball whose boundary is the circumscribed \( (d-1) \)-sphere of \( S_j \). The Delaunay triangulation of \( X \), denoted as \( DT(X) \), is any triangulation of \( X \) such that \( B_j \cap X = \emptyset \) for \( j = 1, \ldots, m \). This is known as the empty-ball property as shown in Figure 1 (a). The Delaunay triangulation generates a mesh of simplices that are most regularized in shape. For point set \( X \subset R^2 \), the Delaunay triangulation \( DT(X) \) maximizes the minimum angle in all the triangles (2-simplices) \( S_1, \ldots, S_m \) over all possible triangulations (Sibson, 1978) as displayed in Figure 1 (b) and (c). As the geometric dual of the Voronoi diagram, the Delaunay triangulation \( DT(X) \) is unique under the assumption that \( X \) is in general position (Delaunay, 1934).

Consider the data \( \{(x_i, y_i) : i = 1, \ldots, n \} \) from model (1), the Delaunay interpolation aims to estimate the conditional expectation function \( E(Y|z) \) for all \( z \in \mathcal{H}(X) \). Generally, there are three steps in the Delaunay interpolation: (i) construct the Delaunay triangulation \( DT(X) \); (ii) find the simplex \( S(z) \in DT(X) \); and (iii) obtain the estimator \( \hat{E}(Y|z) \) by optimizing a target function. For most of Delaunay interpolation methods, the first two steps are identical, while the difference mainly lies in the target function. For example, with \( \gamma_1, \ldots, \gamma_{d+1} \in [0,1] \) such that \( \sum_{k=1}^{d+1} \gamma_k x_{i_k(z)} = z \) and \( \sum_{k=1}^{d+1} \gamma_k = 1 \), the estimator of de Berg et al. (2008) is

\[
\hat{E}(Y|z) = \sum_{k=1}^{d+1} \gamma_k y_{i_k(z)},
\]

which is the minimizer of the squared loss function \( \sum_{i=1}^{n} (y_i - f(x_i))^2 \) among all continuous piecewise linear functions, \( f(z) = \sum_{j=1}^{m} 1_{x \in S_j}(\alpha_j + \beta_j^T z) \). Liu & Yin (2019) introduce a regularization function to balance the model fitting and smoothness of the estimator. However, all the aforementioned approaches require a complete construction of \( DT(X) \), whose size grows exponentially with respect to the dimension \( d \). As a result, no existing algorithms are feasible when the dimension \( d \geq 7 \) due to the limitations of computation time/power and memory space (Chang et al., 2020).

Alternatively, several methods have been proposed for medium- to high-dimensional Delaunay interpolation (Chang et al., 2018ab; 2020). Instead of obtaining the complete \( DT(X) \), these methods only construct the Delaunay simplex \( S(z) \) at each point \( z \) locally and \( E(Y|z) \) can be estimated at a polynomial cost. For any point \( z \in \mathcal{H}(X) \), the DELAUNAYSparse algorithm (Chang et al., 2020) first obtains a seed Delaunay simplex \( S_{\text{seed}} \) close to \( z \). Based on \( S_{\text{seed}} \), Chang et al. (2020) find \( S(z) \) via the breadth first search as described in Algorithm 1 and compute the estimator in (2). Although such an approach is computationally efficient because \( \gamma_1, \ldots, \gamma_{d+1} \) are simultaneously calculated, it only utilizes the information of \( d + 1 \) data points \( \{(x_{i_k(z)}, y_{i_k(z)}) : k = 1, \ldots, d+1 \} \) in the estimation.

\[\text{Algorithm 1 DELAUNAYSparse (Chang et al., 2020)}\]

1: \textbf{Input}: Feature points \( X \), target point \( z \in \mathcal{H}(X) \) and the seed Delaunay simplex \( S_{\text{seed}} \).
2: Let \( S_{\text{current}} = S_{\text{seed}}, A_{\text{frontier}} = \{S_{\text{seed}}\}, A_{\text{explored}} = \emptyset \).
3: \textbf{while} \( z \notin S_{\text{current}} \) \textbf{do}
4: \hspace{8mm} Compute the set of facets of \( S_{\text{current}} \) which is visible to \( z^j \), denoted as \( F_z(S_{\text{current}}) \).
5: \hspace{8mm} \textbf{for} each facet \( F \in F_z(S_{\text{current}}) \) \textbf{do}
6: \hspace{16mm} Grow a new Delaunay simplex \( S_{\text{new}} \neq S_{\text{current}} \) on the facet \( F \) if it exists.
7: \hspace{16mm} \( A_{\text{Frontier}} \leftarrow A_{\text{Frontier}} \cup \{S_{\text{current}}\} \) if \( S_{\text{new}} \) exists and \( S_{\text{new}} \notin A_{\text{explored}} \cup A_{\text{Frontier}} \).
8: \hspace{8mm} \textbf{end for}
9: \hspace{8mm} \( A_{\text{explored}} \leftarrow A_{\text{explored}} \cup \{S_{\text{current}}\} \).
10: \hspace{8mm} \( A_{\text{Frontier}} \leftarrow A_{\text{Frontier}} \setminus \{S_{\text{current}}\} \).
11: \hspace{8mm} \( S_{\text{current}} \leftarrow \) the first simplex in \( A_{\text{Frontier}} \).
12: \textbf{end while}
13: \textbf{Output}: Simplex \( S_{\text{current}} \).

\[\text{A facet } F \text{ of the simplex } S_{\text{current}} \text{ is visible to } z \text{ if there exists a} \]
\[\text{internal point } z' \text{ of } S_{\text{current}} \text{ such that the linear segment from } z \text{ to } z' \text{ intersects } F \text{(Chang et al., 2020)}.\]
This may lead to overfitting and poor estimation when the simplex \( S(z) \) has a small volume and a poorly regularized shape.

### 2.2. Crystallization Search for Delaunay Simplices

As one component of \( DT(\mathcal{X}) \), \( S(z) \) has \( d + 1 \) facets \( F_1, \ldots, F_{d+1} \), each of which is either a facet of \( \mathcal{H}(\mathcal{X}) \) or the shared boundary of \( S(z) \) and one neighbor Delaunay simplex.

**Definition 1.** **Neighbor Delaunay simplices:** Given a set of points \( \mathcal{X} \) and the Delaunay triangulation \( DT(\mathcal{X}) = \{S_1, \ldots, S_m\} \), simplices \( S_j \) and \( S_k \) are neighbors if and only if the intersection \( S_j \cap S_k \) is a shared facet of \( S_j \) and \( S_k \).

Inspired by Algorithm 1 (Chang et al., 2020) which searches \( S(z) \) by growing neighbor Delaunay simplices on the facets of the explored ones recursively, we develop the crystallization search (Algorithm 2) to construct all the Delaunay simplices within the topological distance \( L \) to \( S(z) \), denoted as \( N_L(z) \). Figures 2 and 3 display the crystallization search of \( N_L(z) \) with respect to a target point \( z \in \mathcal{H}(\mathcal{X}) \) and \( L = 0, 1, 2 \) (top row), \( L = 3, 4, 5 \) (bottom row) in \( \mathbb{R}^2 \).

### 2.3. Crystallization Learning

Without loss of generality, let \( V_{z,L} = \cup_{S \in N_L(z)} V(S) \) denote the set of all the data points of the simplices in \( N_L(z) \). Based on the set \( N_L(z) \) of Delaunay simplices topologically closest to the target point \( z \), we propose the crystallization learning to estimate \( E(Y|z) \) by fitting a local linear model, \( E(Y|z) = \alpha + \beta^T z \), to all the data points in \( V_{z,L} \) instead of only the \( d + 1 \) data points of \( S(z) \). Consider that in \( DT(\mathcal{X}) \), a vertex shared by more simplices usually has a larger degree in the network formed by Delaunay edges and thus is more informative in the geometric structure of \( N_L(z) \), we estimate \( \alpha \) and \( \beta \) via the weighted least squares approach

\[
(\hat{\alpha}, \hat{\beta}) = \arg \min_{\alpha, \beta} \sum_{x_i \in V_{z,L}} w_{z,L}(x_i) (y_i - \alpha - \beta^T x_i)^2,
\]
with weight function

\[ w_{z,L}(x) = \left( \sum_{S \in N_L(z)} 1_{\{x \in V(S)\}} \right) \exp \left( -\frac{\|x - z\|_2^2}{m_L(z)} \right), \]

where \( m_L(z) = \left( \sum_{x \in V_{z,L}} \|x - z\|_2 \right) / \left( \sum_{i=1}^{n} 1_{\{x \in V_{x,L}\}} \right). \)

Note that our weight function places more weights on the data points closer to \( z \), similar to the works of Nadaraya (1964) and Watson (1964), and also on those shared by more simplices in \( N_L(z) \). For all \( x \not\in V_{z,L} \), the weights are set as zero. In addition, our weight function is scale-invariant due to the existence of normalization term \( m_L(z) \), i.e., multiplying any constant to features will not change the invariant.

**Theorem 1.** Let \( X \) be a set of \( n \) feature points \( x_1, \ldots, x_n \) in general position and responses \( y_1, \ldots, y_n \) are generated from model (1). The estimator of the crystallization learning, \( E(Y|z) \), is smooth in \( S_k \in DT(\mathbb{X}) \) (\( k = 1, \ldots, m \)).

### 2.4. Selection of \( L \)

Similar to many machine learning methods, the statistical complexity and estimation performance of the crystallization learning is controlled by the hyperparameter \( L \), the maximal topological distance from the generated neighbor Delaunay simplices to \( S(z) \). As a small \( L \) leads to overfitting and a large \( L \) makes \( E(Y|z) \) overly smooth, we propose adapting the leave-one-out cross validation (LOO-CV) to select \( L \) with respect to the target point \( z \) as follows.

1. Compute the Delaunay simplex \( S(z) \) containing \( z \) and values \( \gamma_1, \ldots, \gamma_{d+1} \in [0, 1] \) such that \( \sum_{k=1}^{d+1} \gamma_k x_{i_k}(z) = z \) and \( \sum_{k=1}^{d+1} \gamma_k = 1 \) via Algorithm 1, where \( x_{i_1}(z), \ldots, x_{i_{d+1}}(z) \) are the \( d+1 \) data points of \( S(z) \).
2. For each \( x_{i_k}(z) \in S(z) \), apply the crystallization learning with different candidate values of \( L \) on the leave-one-out data excluding \( (x_{i_k}(z), y_{i_k}(z)) \) to estimate \( y_{i_k}(z) \). Let \( \xi_{i_k}(z,L) \) be the squared estimation error with respect to observation \( (x_{i_k}(z), y_{i_k}(z)) \) and candidate value \( L \).
3. Select the optimal \( L \) as

\[ \tilde{L} = \arg \min_{k=1}^{d+1} \sum_{k=1}^{d+1} \gamma_k \log(\xi_{i_k}(z,L)). \]

### 2.5. Computational Complexity

With vectorized operations in \( \mathbb{R} \), we reduce the average computational complexity of Algorithm 1 from \( \mathcal{O}(d^2n) \) (Chang et al., 2020) to \( \mathcal{O}(d^2 \log n) \). However, as Algorithm 1 is only implemented once, the dominant cost of Algorithm 2 lies in the simplices growth steps (step 4-15). Since in Algorithm 2, the number of generated Delaunay simplices is \( \mathcal{O}(d^2) \) and the average computational complexity of growing a new Delaunay simplex on the facet of \( S_{current} \) is \( \mathcal{O}(\log n) \) with the rank-1 update suggested by Chang et al. (2020), the average computational complexity is \( \mathcal{O}(d^2 \log n) \). Table 1 shows the average runtime in computing \( N_L(z) \) under different scenarios, which validates the \( \mathcal{O}(d^2 \log n) \) complexity of Algorithm 2.

<table>
<thead>
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<th>( L ) ( n ) = 500</th>
<th>( n ) = 1000</th>
<th>( n ) = 2000</th>
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<td>( d = 8 )</td>
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<td>0.11</td>
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<td>( d = 10 )</td>
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<td>( d = 20 )</td>
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<td>( d = 56 )</td>
<td>0.28</td>
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<tr>
<td>( d = 180 )</td>
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<td>( d = 230 )</td>
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<td>( d = 220 )</td>
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</tr>
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<td>( d = 230 )</td>
<td>1.21</td>
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</tr>
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</table>

### 3. Connection with Other Nonparametric Regression Methods

Similar to the two popular paradigms of nonparametric regression methods, i.e., the nearest neighbor and the local linear regression, our crystallization learning consists of three steps in estimating the conditional expectation function \( E(Y|z) \): (i) selecting data points from \( X \) as the neighbors of \( z \) according to a specific criterion; (ii) assigning weights to the selected neighbor data points; and (iii) fitting a local model to the selected neighbor data points.

Since our crystallization learning and the existing methods mainly differ in the first two steps, we compare our crystallization learning with the \( k \)-nearest neighbor (\( k \)-NN) regression and the local linear regression in the computation of neighbor data points. We use the Euclidean distance in the \( k \)-NN regression and the Gaussian kernel in the local linear regression because our crystallization search of neighbor data points is established on the Delaunay triangulation, which is the geometric dual of the Voronoi diagram under the \( L_2 \) norm.

To find the \( k \) nearest neighbor data points, the \( k \)-NN regression computes and sorts the Euclidean distances from the target point \( z \) to all the data points in \( X \). This process can be visualized by the left panel of Figure 4, where a circle with center \( z \) is drawn. The radius keeps increasing until there are \( k \) observed data points falling in or on the circle, which are returned as the \( k \) nearest neighbors. As only the distances are considered, it is likely that the directions from \( z \) to the \( k \) nearest neighbors are not uniformly distributed, especially when \( z \) is close to the boundary of \( \mathcal{H}(X) \) or jump points of the feature data density. The same is true for the local linear regression, where more weights are assigned to the direction from \( z \) to the sample mean. In contrast, the crystallization
Figure 4. Neighbor data points of the target point \( z \) computed by the \( k \)-NN regression with \( k = 5, 10, 15, 20 \) (left panel) and the crystallization learning with \( L = 0, 1, 2, 3 \) (right panel).

Figure 5. The kernel density estimate of the distribution of the directions from the target point \( z \) to its neighbor data points computed by different methods under different values of the hyperparameter. The arrow indicates the direction from the target point \( z \) to the sample mean of \( X \).

Figure 6. The paths of the (weighted) means of neighbor data points computed by different methods as the value of the hyperparameter increases.

Crystallization search computes neighbor data points \( V_{z,L} \) by constructing Delaunay simplices, whose volumes are adaptive to the density of observed data points. As a result, the distances from \( z \) to neighbor data points in \( V_{z,L} \) are different for high-density and low-density directions. This can be seen from Figure 4, where we generate \( x_1, \ldots, x_{100} \in \mathbb{R}^2 \) from the density function,

\[
g(x) \propto \prod_{j=1}^{2}(1 + 0.6 \cdot \text{sign}(x_j)) \exp(-x_j^2/2),
\]

and use different methods to compute the neighbor data points of \( z = (0, 1)^T \). The density function \( g(x) \) is discontinuous at \( z \) with higher density at its right-hand side than its left-hand side. From the left panel of Figure 4, we can see that for all values of \( k \), \( k \)-NN regression computes more neighbor points at the right-hand side of \( z \) than the left-hand side. However, this is not the case for the crystallization learning as exhibited in the right panel of Figure 4. As \( L \) increases, the crystallization learning searches neighbor data points uniformly in all directions, implying the adaptation of our method to the local geometric structure of the data. This can also be observed in Figure 5, where the kernel density estimate (KDE) of the distribution of the directions from the target point \( z \) to its neighbor data points is plotted. The neighbor data points computed by the \( k \)-NN regression and the weights assigned by the local linear regression concentrate in the direction toward the sample mean of \( X \), while the KDE of the crystallization search is much closer to a uniform distribution. As a result, the (weighted) means of neighbor data points under the crystallization search are closer to the target point \( z \) than existing methods as shown in Figure 6.
4. Asymptotic Theory

We first study the asymptotic geometric properties of the Delaunay triangulation $D_T(\mathcal{X})$ under general distribution of feature points and then prove the consistency of the crystallization learning in estimating $E(Y|z)$. All proofs are given in the supplementary materials.

4.1. Asymptotic Geometric Properties

Lemma 1. Let $\mathcal{X}$ be a set of $n$ i.i.d. feature points $x_1, \ldots, x_n$ from a density $g(x)$, which is bounded away from zero and infinity on $\mathbb{R}^d$. Then for any target point $z \in \mathbb{R}^d$, we have $\mathbb{P}(z \in \mathcal{H}(\mathcal{X})) \to 1$ as $n \to \infty$.

By Lemma 1, the target point $z$ falls in $\mathcal{H}(\mathcal{X})$ with asymptotic probability one. Thus, we only consider the inside-hull case in Theorem 2.

Theorem 2. Let $\mathcal{X}$ be a set of $n$ i.i.d. feature points $x_1, \ldots, x_n$ from a density $g(x)$, which is bounded away from zero and infinity on $\mathbb{R}^d$. For any target point $z \in \mathcal{H}(\mathcal{X})$ and any $\rho \in (0, 1)$, we have

$$T(z) = O_p(n^{-\rho/d}),$$

where $T(z) = \max \{||x_i - z||_2: x_i \in \mathbb{V}_{z,L}\}$.

Theorem 2 implies that all the feature points of $\mathbb{V}_{z,L}$ converge to $z$ in probability.

4.2. Consistency

Theorem 3. Let $\mathcal{X}$ be a set of $n$ i.i.d. feature points $x_1, \ldots, x_n$ from a density $g(x)$, which is bounded away from zero and infinity on $\mathbb{R}^d$. Assume that the data $\{(x_i, y_i): i = 1, \ldots, n\}$ are generated from model (1), where the conditional expectation function $E(Y|x)$ is differentiable on $\mathbb{R}^d$, $e_1, \ldots, e_n \in \mathbb{R}$ are i.i.d. random errors with $E(e_i) = 0$ and $E(e_i^2) < \infty$. The estimator obtained from the crystallization learning, $\hat{E}(Y|z)$ with $z \in \mathcal{H}(\mathcal{X})$, satisfies

$$\mathbb{E}\{\hat{E}(Y|z) - E(Y|z)\}^2 \to R_{\min}, \quad \text{as } n \to \infty,$$

where $R_{\min} = \inf_{\hat{f}} \mathbb{E}\{(Y - \hat{f}(x))^2\}$ is the minimal value of the $L_2$ risk over all continuous functions $f : \mathbb{R}^d \to \mathbb{R}$ and $\mathbb{E}$ is the expectation operator with respect to the data $\{x_i, y_i: i = 1, \ldots, n\}$.

5. Experiments

We conduct experiments on synthetic data under different scenarios: (a) to illustrate the effectiveness of our crystallization learning in estimating the conditional mean function $E(Y|z)$; (b) to compare the estimation accuracy of our method with existing nonparametric regression methods, including the $k$-NN regression using the Euclidean distance, the local linear regression using the Gaussian kernel, the multivariate kernel regression using the Gaussian kernel (Hein, 2009) and Gaussian process models; and (c) to validate the proposed data-driven procedure of $L$ selection. We also apply our method to real data to investigate its empirical performance. Given the training data $\{(x_i, y_i): i = 1, \ldots, n\}$, we use the mean squared error (MSE) under the method $M$, $\text{MSE}_M = \frac{1}{n} \sum_{k=1}^{n} (\hat{E}_M(Y|x_k) - E(Y|x_k))^2$, to evaluate the accuracy of the estimator $\hat{E}_M(Y|z)$ at random target points $z_1, \ldots, z_{100} \in \mathcal{H}(\mathcal{X})$.

5.1. Experiments on Synthetic Data

In the experiments on synthetic data, we consider two scenarios to investigate the performance of our crystallization learning at general internal points of $\mathcal{H}(\mathcal{X})$ and jump points of the data density. For each scenario, we simulate 100 training datasets $\{(x_i, y_i): i = 1, \ldots, n\}$ and corresponding sets of target points $\{z_1, \ldots, z_{100}\}$ under different values of $n$ and $d$ as follows.

- **Scenario 1:** (General internal points) For each dataset, $x_1, \ldots, x_n$ are independently sampled from the multivariate normal distribution $\text{MVN}(0, \mathbf{I}_d)$ with an identity covariance matrix $\mathbf{I}_d$. The responses $y_1, \ldots, y_n$ are generated from an additive model,

$$Y|x \sim N\left(\sum_{j=1}^{d} c_j g_j(x_j), 1\right),$$  

where $x = (x_1, \ldots, x_d)^T$, $c_1, \ldots, c_d \sim N(0, 1)$, $g_j(\cdot) = \sum_{k=1}^{10} b_{jk} \phi(\cdot; \mu_{jk}, \sigma_{jk}^2)$, $b_{jk} \sim N(0, 1)$, $\mu_{jk} \sim N(0, 1)$, $\sigma_{jk}^2 \sim \text{Gamma}(1, 1)$, and $\phi(\cdot; \mu_{jk}, \sigma_{jk}^2)$ is the density of $N(\mu_{jk}, \sigma_{jk}^2)$, $j = 1, \ldots, d; k = 1, \ldots, 10$. For $k = 1, \ldots, 100$, target point $z_k$ is generated as $z_k = \sum_{i=1}^{n} \omega_{ik} x_i$, with $(\omega_{1k}, \ldots, \omega_{nk}) \sim \text{Dirichlet}(1_n)$.

- **Scenario 2:** (Jump points of the data density) For each dataset, $x_1, \ldots, x_n$ are sampled from the distribution,

$$g(x) = 2^{-d} \prod_{j=1}^{d} (1 + 0.4 \cdot \text{sign}(x_j)) \exp(-|x_j|),$$

and responses $y_1, \ldots, y_n$ are generated from the additive model (4). It is clear that the density $g(x)$ jumps at the point set $\{x \in \mathbb{R}^d: \prod_{j=1}^{d} x_j = 0\}$. For $k = 1, \ldots, 100$, target point $z_k$ is generated as $z_k = \sum_{i=1}^{n} \omega_{ik} x_i \otimes s_k$, where $(\omega_{1k}, \ldots, \omega_{nk}) \sim \text{Dirichlet}(1_n)$, $\otimes$ is the elementwise multiplication operator, $s_k = (s_{k1}, \ldots, s_{kj})^T$ and $s_{k1}, \ldots, s_{kj} \sim \text{Bernoulli}(0.7)$. 
Table 2. The averaged values of \( \log(\text{MSE}) \) and standard deviations in parentheses using crystallization learning (CL) in comparison with \( k\text{-NN} \) \((k = 5, 10, k^*)\), where \( k^* \) equals the size of \( \mathcal{V}_{x,L} \), local linear (LL) regression, kernel regression (KR) and Gaussian process (GP) in estimating \( E(Y|z) \) under different scenarios, different sample sizes \((n)\) and different dimensions of the feature space \((d)\).

<table>
<thead>
<tr>
<th>(d)</th>
<th>(n)</th>
<th>(\log(\text{MSE}_{CL}))</th>
<th>(\log(\text{MSE}_{\text{5-NN}}))</th>
<th>(\log(\text{MSE}_{\text{LL}}))</th>
<th>(\log(\text{MSE}_{\text{KR}}))</th>
<th>(\log(\text{MSE}_{\text{GP}}))</th>
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<td>5</td>
<td>200</td>
<td>-1.11(0.21)</td>
<td>0.23(0.09)</td>
<td>0.12(0.09)</td>
<td>0.33(0.11)</td>
<td>0.56(0.11)</td>
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<td>0.37(0.11)</td>
<td>0.45(0.13)</td>
<td>0.91(0.17)</td>
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<td>0.53(0.13)</td>
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<td>1.18(0.19)</td>
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<tr>
<td></td>
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<td>0.48(0.14)</td>
<td>0.38(0.14)</td>
<td>0.59(0.16)</td>
<td>1.06(0.22)</td>
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<td>0.13(0.07)</td>
<td>0.14(0.08)</td>
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<td>0.43(0.13)</td>
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<td></td>
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<td>0.51(0.11)</td>
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<tr>
<td>20</td>
<td>200</td>
<td>1.46(0.14)</td>
<td>0.14(0.08)</td>
<td>-0.02(0.06)</td>
<td>-0.01(0.06)</td>
<td>-0.02(0.03)</td>
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<tr>
<td></td>
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<td>0.25(0.10)</td>
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<td>0.36(0.10)</td>
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<tr>
<td></td>
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<td>0.18(0.10)</td>
<td>-0.01(0.06)</td>
<td>0.02(0.04)</td>
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</table>

In both scenarios, we apply our crystallization learning and existing methods to estimate \( E(Y|z) = \sum_{j=1}^{d} c_j g_j(z) \) at target points \( z_1, \ldots, z_{100} \). We implement the crystallization learning with \( L = 3 \) (or \( L = 2 \)) when \( d = 5 \) and 10 (or \( d = 20 \) and 50) and obtain the estimator \( \hat{E}(Y|z) \). We implement the \( k\text{-NN} \) regression with \( k = 5, 10, k^* \), where \( k^* \) equals the size of \( \mathcal{V}_{x,L} \), and the local linear regression and kernel regression with bandwidth \( h = 1 \).

The estimation results of our method averaged over 100 simulations under different scenarios and different values of \( n \) and \( d \) are displayed in Table 2. For each \( d \), the estimation accuracy of the crystallization learning improves as the sample size increases, indicating its consistency in estimating the conditional mean function \( E(Y|z) \) in the convex hull \( \mathcal{H}(X) \). For low-dimensional cases \((d = 5 \text{ or } 10)\), the crystallization learning generally outperforms the existing methods as the sample size grows in both scenarios, demonstrating that our method is more efficient asymptotically. In scenario 1, although our method does not completely dominate the existing ones when \( d = 20 \) or 50, its performance is still superior or at least comparable to the existing methods. In the high-dimensional cases \((d = 20 \text{ or } 50)\) of scenario 2, the crystallization learning is not dominated by any existing approaches, suggesting the robustness of our method to the variation or sudden change of the data density. Overall, our crystallization learning performs well and is stable in estimating the conditional mean function \( E(Y|z) \) at general internal points of \( \mathcal{H}(X) \) and jump points of the data density.

5.2. Experiments on Real Data

To illustrate the empirical performance of the crystallization learning, we apply it to several real datasets from the UCI repository. The first one is the critical assessment of
Figure 7. Boxplots of $\log(\text{MPSE}_M/\text{MPSE}_{CL})$ corresponding to $k$-NN ($k = 5, 10, k^*$, where $k^*$ equals the size of $V_{k,L}$), local linear (LL) regression, kernel regression (KR) and Gaussian process (GP) in estimating $E(Y|z)$ under different datasets and sizes of the training set ($n$).

The protein structure prediction (CASP) dataset\(^2\) (Betancourt & Skolnick, 2001) of experimental records on protein structure prediction. The CASP dataset includes 45730 records of 9 features, where the response is the root mean squared deviation (RMSD) of the residues. The second is the Concrete dataset\(^3\) (Yeh, 1998), which consists of 1030 experimental records of concrete compressive strength measurement. We use the content of 7 concrete ingredients and the age of a concrete sample to predict its compressive strength. The last one is Parkinson’s telemonitoring dataset\(^4\) (Tsanas et al., 2010), which is composed of 5875 voice recordings of 16 biomedical voice measures from 42 patients with early-stage Parkinson’s disease in a six-month trial. We use these 16 biomedical voice measures to predict the motor and total UPDRS (unified Parkinson’s disease rating scale) scores.

For each dataset, we take 100 bootstrap samples without replacement of size $n$ ($n = 200, 500, 1000$ or $2000$) for training and 100 samples of size 100 for testing to compare the crystallization learning with existing methods. To eliminate the impact of feature correlations and scales, we extract and standardize the principal components of features in the training set as the observed feature points $x_1, \ldots, x_n$, and the same transformation is applied to the testing set to obtain the target points $z_1, \ldots, z_{100}$. We take $L = 3$ for crystallization learning, implement the $k$-NN regression with $k = 5, 10, k^*$, where $k^*$ equals the size of $V_{k,L}$, and the local linear regression and kernel regression with bandwidth $h = 1$. As the true values of conditional expectations $E(Y|z_1), \ldots, E(Y|z_{100})$ are unknown in real application, we quantify the performance of the
method $\mathcal{M}$ by the mean predictive squared error (MPSE): 
$$\text{MPSE}_{\mathcal{M}} = \frac{1}{100} \sum_{k=1}^{100} \left( \hat{E}_{\mathcal{M}}(Y|x_k) - y_k \right)^2.$$ 

Figure 7 shows the comparison results averaged over 100 bootstrap samples between our method and existing ones under different datasets and sizes of the training set ($n$). For each dataset, it is clear that as $n$ increases, the advantage of the crystallization learning over most of existing methods amplifies. Overall, the crystallization learning dominates all the existing methods in most of the cases.

### 5.3. Experiments on $L$ Selection

To validate the data-driven procedure of $L$ selection proposed in Section 2.4, we conduct experiments under Scenario 1 of Section 5.1. With candidate values $1, \ldots, 8$ of $L$ and $d = 5$, 100 training datasets and corresponding sets of target points are simulated for different sample sizes $n$ ($n = 200, 500, 1000, 2000$) respectively to investigate whether the proposed procedure improves the estimation accuracy of the crystallization learning.

The estimation results of our method with different candidate values of $L$ and the selected $\hat{L}$ averaged over 100 simulations under different sample sizes are displayed in Figure 8. It is clear that as the sample size $n$ increases, the optimal $L$ with smallest averaged value of $\log(\text{MSE}_L)$ grows. This is reasonable because that a larger $n$ would lead to smaller volumes of simplices in $\mathcal{N}_L(z)$ and thus a larger $L$ is needed for accurate estimation. In addition, the averaged value of $\log(\text{MSE}_L)$ is closer to the smallest averaged value of $\log(\text{MSE}_L)$ when $n$ is larger, suggesting the effectiveness of our LOO-CV procedure in improving the estimation accuracy.

### 6. Conclusions

The Delaunay triangulation is a powerful tool to partition the feature space in a data-driven way, which has the least roughness for smooth surface reconstruction. We incorporate the Delaunay triangulation into the framework of nonparametric regression and develop the crystallization learning procedure. Without the need to triangulate the entire feature space which becomes infeasible for high-dimensional cases, our method conducts the Delaunay triangulation locally at each specific target point like crystal growth. The conditional expectation $E(Y|x)$ at the target point $x \in \mathcal{H}(\mathcal{Z})$ is estimated by fitting a linear model to the data points of the Delaunay simplices computed by the crystallization search. Compared with existing nonparametric regression methods, our method is more adaptive to the local geometric structure of the data, which computes the neighbor data points uniformly in all directions and their weighted mean is closer to the target point $z$. Both theoretical studies and numerical experiments show that the crystallization learning is consistent in estimating $E(Y|x)$ and it generally outperforms the existing methods.

There are still some directions in which we can further develop our work. Given that the crystallization learning is a local approach, it is possible to combine it with the uniform design and develop a global version of crystallization learning in a hierarchical way. As our method searches $\mathcal{N}_{z,L}$ in a deterministic way, we can also develop the stochastic crystallization search to reduce the boundary effect on the estimator $\hat{E}(Y|x)$. Other possible extensions include extrapolation of $E(Y|x)$ at $z \notin \mathcal{H}(\mathcal{Z})$ with the Möbius transformation (Zhou et al., 2019), regression problems in other metric spaces (e.g., manifold regression and structured output) as discussed in Hein (2009) and online regression problems (Kuzborskij & Cesa-Bianchi, 2017).

### Acknowledgement

We thank the four anonymous reviewers for insightful suggestions that have significantly improved the paper. This research is supported by the Research Grants Council of Hong Kong (17326316) and TCL Corporate Research (Hong Kong).
References


